

(1)

Spring 2014

Multidimensional Scaling (MDS)

Note Title

11/2/2010

MDS is a linear projection method. It is related to PCA. MDS and PCA can be used for non-linear projection - next lecture.

Key Idea of MDS: project to preserve the distances $|\underline{x}_i - \underline{x}_j|$ between the data points $\{\underline{x}_i : i=1..n\}$, i.e. $\underline{x}_i \rightarrow \underline{y}_i$ such that $|\underline{x}_i - \underline{x}_j| \approx |\underline{y}_i - \underline{y}_j|$, but the \underline{y} 's have lower dimension.

This projection constraint is imposed on the dot products $\underline{x}_i \cdot \underline{x}_j \approx \underline{y}_i \cdot \underline{y}_j$. This will imply that $|\underline{x}_i - \underline{x}_j| \approx |\underline{y}_i - \underline{y}_j|$.

Important Property: we only need to know $|\underline{x}_i - \underline{x}_j| \triangleq \Delta_{ij}$ in order to calculate $\underline{x}_i \cdot \underline{x}_j$. This will be useful for non-linear applications later. Also, sometimes only $|\underline{x}_i - \underline{x}_j|$ is specified.

Result: $\underline{x}_i \cdot \underline{x}_j = \frac{1}{ZN} \sum_k \Delta_{ik}^2 + \frac{1}{ZN} \sum_i \Delta_{ij}^2 - \frac{1}{ZN^2} \sum_k \Delta_{ik}^2 - \frac{1}{Z} \Delta_{ij}^2$
 provided $\sum_i \underline{x}_i = 0$. (Subtract $\frac{1}{N} \sum_i \underline{x}_i$ from data to ensure this.)

Proof: $\Delta_{ij}^2 = |\underline{x}_i|^2 + |\underline{x}_j|^2 - 2 \underline{x}_i \cdot \underline{x}_j$

Let $T = \sum_i |\underline{x}_i|^2$, note that $\sum_i \underline{x}_i \cdot \underline{x}_j = 0 = \sum_j \underline{x}_i \cdot \underline{x}_j$.

Then $\sum_k \Delta_{ik}^2 = N |\underline{x}_i|^2 + T$, $\sum_i \Delta_{ij}^2 = N |\underline{x}_j|^2 + T$, $\sum_k \Delta_{ik}^2 = ZNT$.

The result follows by substitution.

Define the Gram matrix

$$G_{ij} = \underline{x}_i \cdot \underline{x}_j = \frac{1}{Z} \sum_k \Delta_{ik}^2 + \frac{1}{Z} \sum_i \Delta_{ij}^2 - \frac{1}{Z} \sum_k \Delta_{ik}^2 - \frac{1}{Z} \Delta_{ij}^2.$$

Define an error function

$$\text{err}(\underline{y}) = \sum_{ij} (G_{ij} - \underline{y}_i \cdot \underline{y}_j)^2$$

\underline{x}_i lies in D-dim space, G_{ij} is $N \times N$ matrix

\underline{y}_i is a vector in d-dim space $d \ll D$

$$d < N$$

(2)

$$\text{Minimize } \text{err}(\underline{y}) = \sum_{i,j} (G_{ij} - \underline{y}_i \cdot \underline{y}_j)^2.$$

Do spectral decomposition:

$$G = \sum_{d=1}^N \lambda_d \underline{v}_d \underline{v}_d^T$$

$$\lambda_1 > \dots > \lambda_N > 0$$

eigenvalues of G

$$\underline{v}_1, \underline{v}_2, \dots, \underline{v}_N$$

eigenvectors

Claim: optimal minimization

$$\text{is } \underline{y}_d = \sqrt{\lambda_d} \underline{v}_d$$

$$\text{ie. } \underline{y}_i = (\sqrt{\lambda_1} v_i^1, \sqrt{\lambda_2} v_i^2, \dots, \sqrt{\lambda_N} v_i^N)$$

N -dimension

$$\text{Proof. } \underline{y}_i \cdot \underline{y}_j = \sum_d \sqrt{\lambda_d} v_i^d \sqrt{\lambda_d} v_j^d = \sum_d \lambda_d v_i^d v_j^d = G_{ij}$$

This gives $\text{err} = 0$.

We can reduce the dimension by

truncating \underline{y}_i to $\underline{y}_i = (\sqrt{\lambda_1} v_i^1, \dots, \sqrt{\lambda_d} v_i^d)$
for $d < N$

In this case $G_{ij} \neq \underline{y}_i \cdot \underline{y}_j$

$$\underline{y}_i \cdot \underline{y}_j = \sum_{d=1}^d \lambda_d v_i^d v_j^d, \quad G_{ij} = \sum_{d=1}^N \lambda_d v_i^d v_j^d$$

$$\text{Hence } G_{ij} - \underline{y}_i \cdot \underline{y}_j = \sum_{d=d+1}^N \lambda_d v_i^d v_j^d$$

$$\text{Claim } \sum_{i,j} (G_{ij} - \underline{y}_i \cdot \underline{y}_j)^2 = \sum_{d=d+1}^N \lambda_d^2$$

$$\text{Proof. } \sum_{i,j} \left\{ \sum_{d=d+1}^N \sum_{\beta=d+1}^N \lambda_d \lambda_\beta v_i^d v_j^d v_i^\beta v_j^\beta \right\}$$

$$\sum_{d=d+1}^N \sum_{\beta=d+1}^N \lambda_d \lambda_\beta v_i^d v_j^\beta = \sum_i v_i^d v_i^\beta = S^{d\beta}, \quad \sum_j v_j^d v_j^\beta = S^{d\beta}$$

$$\sum_{d=d+1}^N \sum_{\beta=d+1}^N \lambda_d \lambda_\beta S^{d\beta} S^{\beta d} = \sum_{d=d+1}^N \lambda_d^2.$$

Hence we project to d -dimension

provided $\sum_{d=d+1}^N \lambda_d^2$ is small, or $\frac{\sum_{d=d+1}^N \lambda_d^2}{\sum_{d=1}^N \lambda_d^2}$ is small.

$$\underline{y}_i = (\sqrt{\lambda_1} v_i^1, \dots, \sqrt{\lambda_d} v_i^d)$$

(3) Relation between MDS & PCA?
 Both linear. Both depend on eigenvectors/eigenvalues.

Recall PCA (subtract mean to ensure $\sum_i \underline{x}_i = 0$)

$$\underline{X}_p = (\underline{\underline{x}}_1, \dots, \underline{\underline{x}}_d)$$

the $\underline{\underline{e}}$'s are eigenvectors of $K_{ab} = \frac{1}{N} \sum_{i=1}^N \underline{\underline{x}}_i \underline{\underline{x}}_i^T$

$$\text{MDS } \underline{\underline{y}}_i = (\sqrt{\lambda_1} v_1^1, \dots, \sqrt{\lambda_d} v_d^1)$$

the $\underline{\underline{v}}$'s are eigenvectors of $G_{ij} = \sum_{i=1}^N \underline{\underline{x}}_i \underline{\underline{x}}_j$

Claim: the eigenvalues of $\underline{\underline{G}}$ and $\underline{\underline{K}}$ are the same. The eigenvectors are closely related.

Proof Let $\underline{\underline{X}}$ be an $N \times D$ matrix $i = 1 \text{ to } N$ no. of points
 with elements X_{ia} $a = 1 \text{ to } D$ space dimensions
 $\underline{\underline{X}}$ a^{th} component of i^{th} datapoint

Consider $\underline{\underline{X}} \underline{\underline{X}}^T$ $N \times N$ matrix, $(\underline{\underline{X}} \underline{\underline{X}}^T)_{ij} = \sum_a X_{ia} X_{ja}$

$\underline{\underline{X}}^T \underline{\underline{X}}$ $D \times D$ matrix, $(\underline{\underline{X}}^T \underline{\underline{X}})_{ab} = \sum_i X_{ia} X_{ib}$

Both are square matrices and both are positive definite, so they have positive eigenvalues.

$\underline{\underline{X}} \underline{\underline{X}}^T$ is used for MDS, $\underline{\underline{X}}^T \underline{\underline{X}}$ is used for PCA

Suppose $\underline{\underline{e}}, \lambda$ are an eigenvector, eigenvalue of $\underline{\underline{X}}^T \underline{\underline{X}}$

$$\underline{\underline{X}}^T \underline{\underline{X}} \underline{\underline{e}} = \lambda \underline{\underline{e}}$$

$$\text{so } \underline{\underline{X}} \underline{\underline{X}}^T \underline{\underline{X}} \underline{\underline{e}} = \lambda \underline{\underline{X}} \underline{\underline{e}}$$

$$(\underline{\underline{X}} \underline{\underline{X}}^T) (\underline{\underline{X}} \underline{\underline{e}}) = \lambda (\underline{\underline{X}} \underline{\underline{e}})$$

so $(\underline{\underline{X}} \underline{\underline{e}})$ is an eigenvector (un-normalized)
 of $\underline{\underline{X}} \underline{\underline{X}}^T$ with eigenvalue λ .

Similarly if $\underline{\underline{X}} \underline{\underline{X}}^T \underline{\underline{v}} = \lambda \underline{\underline{v}}$ then $(\underline{\underline{X}}^T \underline{\underline{v}})$ is an eigenvector (un-normalized)
 of $\underline{\underline{X}}^T \underline{\underline{X}}$ with eigenvalue λ .

Conclusion \rightarrow the two matrices have the same eigenvalues
 and related eigenvectors.

(4)

Result \rightarrow the truncation conditions for MDS
and PCA are similar $\sum_{i=1}^k \lambda_i^2 > \text{Threshold}$.

MDS projects $\underline{\underline{X}}$ to $\underline{\underline{y}}_i = (\sqrt{\lambda_1} v_1^1, \dots, \sqrt{\lambda_k} v_k^1)$

PCA projects $\underline{\underline{X}}$ to $\underline{\underline{y}} = (\underline{\underline{x}} \cdot \underline{e}_1, \dots, \underline{\underline{x}} \cdot \underline{e}_d)$

where the \underline{e} 's and the \underline{v} 's are related
(see previous page).

Note: The equivalence between the eigenvalues

of $\underline{\underline{X}} \underline{\underline{X}}^T$ and $\underline{\underline{X}}^T \underline{\underline{X}}$ has computational importance

If $N \ll D$, faster to compute eigenvalues/vectors
for $(\underline{\underline{X}} \underline{\underline{X}}^T)$, then convert to eigenvalues/vectors of $\underline{\underline{X}}^T \underline{\underline{X}}$

Note: to do PCA we have to know the data $\{\underline{\underline{x}}\}$

but to do MDS we only need to know Δ_{ij} .

In some applications it is possible to specify Δ_{ij}
but not the $\{\underline{\underline{x}}_i\}$.

Deeper Understanding: This relationship between $\underline{\underline{X}} \underline{\underline{X}}^T$ and $\underline{\underline{X}}^T \underline{\underline{X}}$ can
be used to prove SVD:

$$\underline{\underline{X}} = \underline{\underline{F}} \underline{\underline{D}} \underline{\underline{E}} \quad \text{where } \underline{\underline{F}} \underline{\underline{F}}^T = \underline{\underline{I}} \quad \underline{\underline{E}} \underline{\underline{E}}^T = \underline{\underline{I}}$$

$$\underline{\underline{D}} = \begin{pmatrix} d_1 & 0 \\ 0 & d_0 \end{pmatrix}$$

This is a generalization of the spectral decomposition

$\underline{\underline{G}} = \sum \lambda_a \underline{\underline{V}}_a \underline{\underline{V}}_a^T$ $\underline{\underline{X}}$ is $N \times D$ $N \neq D$
to any matrix $\underline{\underline{X}}$ \rightarrow so $\underline{\underline{X}}$ is not square.

It follows that $\underline{\underline{X}}^T \underline{\underline{X}} = (\underline{\underline{E}}^T \underline{\underline{D}} \underline{\underline{F}}^T) (\underline{\underline{F}} \underline{\underline{D}} \underline{\underline{E}})$

$$= \underline{\underline{E}}^T \underline{\underline{D}}^2 \underline{\underline{E}}$$

spectral decomposition \rightarrow with $\lambda_1 = d_1^2, \lambda_2 = d_2^2 \dots$

$$\text{and } \underline{\underline{X}}^T \underline{\underline{X}} = \underline{\underline{F}} \underline{\underline{D}} \underline{\underline{E}} \underline{\underline{E}}^T \underline{\underline{D}}^2 \underline{\underline{F}}^T = \underline{\underline{F}} \underline{\underline{D}}^2 \underline{\underline{F}}^T$$

So SVD is like the square root of spectral decomposition!